

## VECTOR AUTOREGRESSIVE (VAR) MODELS

It is quite common in economics to have models in which some variables are not only explanatory variables for a given dependent variables, but are also explained by the variables that they are used to determine.

In these cases, we have models of simultaneous equations, in which it is necessary to identify endogenous and exogenous variables.

The decision regarding such a differentiation among variables was heavily criticized by Sims (1980). According to Sims, if there is simultaneity among a number of variables, then all <sup>these</sup> variables should be treated in the same way.

In other words, there should be no distinction between endogenous and exogenous variables.

Therefore, once this distinction is abandoned, all variables are treated as endogenous. This means that in its general reduced form each equation has the set of regressors, which leads to the development of VAR models.

## The VAR MODEL

When we are not confident that a variable really is exogenous, each variable has to be treated symmetrically.

Take, for example, the time series  $y_t$  that is affected by current and past values of  $x_t$  and; simultaneously, the time series  $x_t$  to be a series that is affected by current and past values of the  $y_t$  series



$$(1) \quad Y_t = \beta_{10} + \beta_{12} X_t + \gamma_{11} Y_{t-1} + \gamma_{12} X_{t-1} + u_{1t}$$

$$(2) \quad X_t = \beta_{20} - \beta_{21} Y_t + \gamma_{21} Y_{t-1} + \gamma_{22} X_{t-1} + u_{2t}$$

where we assume that both  $Y_t$  and  $X_t$  are stationary [we will see this issue later] and  $u_{1t}$  and  $u_{2t}$  are uncorrelated white-noise error terms.  $u_{2t}$  and  $u_{1t}$  are also called structural shocks.

Equations (1) and (2) constitute a first-order VAR model, because the longest lag is unity.

These equations are not reduced-form equations, since  $Y_t$  has a contemporaneous impact on  $X_t$  (given by  $-\beta_{21}$ ) and  $X_t$  has a

contemporaneous impact on  $Y_t$  (given by  $\beta_{12}$ ).

Rewriting the system using matrix algebra, we get

$$\begin{cases} Y_t + \beta_{12} X_t = \beta_{10} + \gamma_{11} Y_{t-1} + \gamma_{12} X_{t-1} + u_{1t} \\ \beta_{21} Y_t + X_t = \beta_{20} + \gamma_{21} Y_{t-1} + \gamma_{22} X_{t-1} + u_{2t} \end{cases}$$

→ Structural VAR (SVAR) or the primitive system

$$(3) \quad \underbrace{\begin{bmatrix} 1 & \beta_{12} \\ \beta_{21} & 1 \end{bmatrix}}_{\underline{B}} \underbrace{\begin{bmatrix} y_t \\ x_t \end{bmatrix}}_{\underline{z}_t} = \underbrace{\begin{bmatrix} \beta_{10} \\ \beta_{20} \end{bmatrix}}_{\underline{\theta}_0} + \underbrace{\begin{bmatrix} \gamma_{11} & \gamma_{12} \\ \gamma_{21} & \gamma_{22} \end{bmatrix}}_{\underline{\theta}_1} \underbrace{\begin{bmatrix} y_{t-1} \\ x_{t-1} \end{bmatrix}}_{\underline{z}_{t-1}} + \underbrace{\begin{bmatrix} u_{yt} \\ u_{xt} \end{bmatrix}}_{\underline{u}_t}$$

$$(4) \quad \underline{B} \underline{z}_t = \underline{\theta}_0 + \underline{\theta}_1 \underline{z}_{t-1} + \underline{u}_t \quad \left. \vphantom{\underline{B} \underline{z}_t} \right\} \begin{array}{l} \text{Structural VAR} \\ \text{(SVAR) or the} \\ \text{Primitive system} \end{array}$$

Multiplying both sides by  $\underline{B}^{-1}$  yields

$$\underline{z}_t = \underline{B}^{-1} \underline{\theta}_0 + \underline{B}^{-1} \underline{\theta}_1 \underline{z}_{t-1} + \underline{B}^{-1} \underline{u}_t$$

$$(5) \quad \left( \begin{array}{l} \text{VAR in} \\ \text{standard form} \\ \text{unstructured} \\ \text{VAR} = \text{UVAR} \end{array} \right) \boxed{\underline{z}_t = \underline{A}_0 + \underline{A}_1 \cdot \underline{z}_{t-1} + \underline{e}_t} \quad \left. \vphantom{\underline{z}_t} \right\} \begin{array}{l} \text{Reduced} \\ \text{form} \end{array}$$

For purposes of notational simplification we can denote as  $a_{i0}$  the  $i^{\text{th}}$  element of the vector  $\underline{A}_0$ ;  $a_{ij}$  the element in row  $i$  and column  $j$  of the matrix  $\underline{A}_1$ ; and  $e_{it}$  as the  $i^{\text{th}}$  element of the vector  $\underline{e}_t$

Note: To decide the # of lags  
we can use Akaike or Schwarz criteria  
and choose with lowest ones.

Using this notation, we can rewrite the

VAR model as;

$$(6) \quad Y_t = a_{10} + a_{11} Y_{t-1} + a_{12} X_{t-1} + e_{1t}$$
$$(7) \quad X_t = a_{20} + a_{21} Y_{t-1} + a_{22} X_{t-1} + e_{2t}$$

impulses  
or  
innovations  
or  
shocks

Here the new error terms,  $e_{1t}$  and  $e_{2t}$ , are  
composites of the two shocks  $u_{Yt}$  and  $u_{Xt}$ .  
 $e_{1t}$  and  $e_{2t}$  are also white noise processes. ~~⊗~~

Pros and Cons of the VAR Models

The VAR model approach has some very  
good characteristics. First, it is very simple.

(1) The econometrician does not have to worry  
about which variables are endogenous  
or exogenous.

(2) Estimation is also very simple, in the  
sense that each equation can be  
estimated separately with the usual

OLS method.

→

(3) Forecasts obtained from VAR models are in most cases better than those obtained from far more complex simultaneous equation models

However, VAR models have also faced severe criticism over various points

- (1) Firstly, they are atheoretic, in that they are not based on any economic theory. In addition, "everything causes everything."
- (2) Another problem is the loss of degrees of freedom. If we suppose that we have a 3 variable VAR model and decide to include 12 lags from each variable in each equation, the degrees of freedom will decline by 36 plus 1 (for intercept), i.e., by 37. If the sample size is not sufficiently large, estimating will be very problematic.
- (3) The obtained coefficients of the VAR model are difficult to interpret because of their lack of any theoretical background.

To overcome this criticism, the advocates of VAR models estimate so-called impulse response functions. The impulse response function examines the response of the dependent variable in the VAR to shocks in the error terms. The difficult issue here, however, is defining the shocks.

The general view is that we would like to shock the structural errors ( $u_{yt}$  and  $u_{xt}$ ) which we can interpret easily as a shock to a particular part of a structural model.

However, we only observe the reduced-form errors in Equations (6) and (7), and these are made up of a combination of structural errors

$$e_{1t} = \frac{u_{yt} + \beta_{12}u_{xt}}{1 - \beta_{12}\beta_{21}}$$

$$e_{2t} = \frac{u_{xt} + \beta_{21}u_{yt}}{1 - \beta_{12}\beta_{21}}$$

So we have to disentangle the structural errors in some way, and this is known as the identification problem. There are a variety of ways of doing this, however the different methods can give rise to quite different results and there is no objective statistical criteria for choosing between these different methods.

## Causality Tests

Causality in econometrics is somewhat different from the concept in everyday use; it refers more to the ability of one variable to predict (and therefore cause) the other. The relationship between these variables can be captured by a VAR model. In this case it is possible to state that

- (1)  $y_t$  causes  $x_t$
- (2)  $x_t$  causes  $y_t$
- (3) There is a bi-directional feedback
- (4) Two variables are independent



Granger (1969) developed a relatively simple test that defined causality as follows:

"a variable  $Y_t$  is said to be Granger cause  $X_t$  if  $X_t$  can be predicted with greater accuracy by using past values of the  $Y_t$  variable rather than not using such past values, all other terms remaining unchanged."

## The Granger Causality Test

The Granger causality test for the case of two stationary variables  $Y_t$  and  $X_t$  involves as a first step the estimation of the following VAR model

$$Y_t = a_1 + \sum_{i=1}^n \beta_i X_{t-i} + \sum_{j=1}^m \gamma_j Y_{t-j} + e_{1t} \quad (1)$$

$\beta_i$  sig.  $\rightarrow$  not sig.  $\rightarrow$  X granger causes Y

$$X_t = a_2 + \sum_{i=1}^n \theta_i X_{t-i} + \sum_{j=1}^m \delta_j Y_{t-j} + e_{2t} \quad (2)$$

$\theta_i$  not sig.  $\delta_j$  signif.  $\rightarrow$  Y Granger cause X

Where it is assumed that both  $u_{1t}$  and  $u_{2t}$  are uncorrelated white-noise error terms.

In this model we can have the following different cases:

Case 1 The lagged  $X_t$  terms in equation (1) may be statistically different from zero as a group, and the lagged  $Y_t$  terms in equation (2) not statistically different from zero. In this case we see  $X_t$  <sup>(Granger)</sup> causes  $Y_t$ .

Case 2 The lagged  $Y_t$  terms in Eq (2) may be statistically different from zero as a group, and the lagged  $X_t$  terms in Eq (1) not statistically different from zero. In this case we see that  $Y_t$  <sup>(Granger)</sup> causes  $X_t$ .

Case 3 Both sets of  $X_t$  and  $Y_t$  terms are statistically different from zero in Eq (1) and (2), so that there is bi-directional causality.

Case 4 Both sets of  $X_t$  and  $Y_t$  terms are not statistically different from zero in Eq (1) and (2), so that  $X_t$  is independent of  $Y_t$ .

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The Granger causality test, then involves the following procedure. First, estimate the VAR model given by the Equations (1) and (2). Then check the significance of the coefficients and apply variable deletion tests, first in the lagged  $X$  terms for Equation (1) and then in the lagged  $Y$  terms for Eq. (2).

Steps are as follows

Step 1 Regress  $Y_t$  on lagged  $Y$  terms as follows:

$$Y_t = a_1 + \sum_{j=1}^m \gamma_j Y_{t-j} + e_{1t} \quad (3)$$

and obtain the SSR of this regression (restricted model) and label it as  $SSR_R$ .

Step 2 Regress  $Y_t$  on lagged  $Y$  terms plus lagged  $X$  terms as in Equation (1)

$$Y_t = a_1 + \sum_{i=1}^n \beta_i X_{t-i} + \sum_{j=1}^m \gamma_j Y_{t-j} + e_{1t} \quad (1)$$

and obtain the SSR of this regression (unrestricted model) and label it as  $SSR_u$ .

Step 3 Set the  $H_0$  and  $H_A$  as:

$$H_0: \sum_{i=1}^n \beta_i = 0 \quad (\text{or } X_t \text{ does not cause } Y_t)$$

$$H_A: \sum_{i=1}^n \beta_i \neq 0 \quad (\text{or } X_t \text{ does cause } Y_t)$$

Step 4 calculate the  $Q$  statistic

$$Q = \frac{(SSR_R - SSR_U) / P}{SSR_U / (T - k - 1)}$$

where  $k = m + n$  and  $p$  is the # of restrictions (so  $p = n$ )

$$Q \sim F_{P, T-k-1}^{\alpha}$$

Step 5 If  $Q > F_{P, T-k-1}^{\alpha} \Rightarrow R' H_0$  and conclude that

$X$  Granger causes  $Y$

# GRANGER CAUSALITY

- We have to be aware of that Granger causality does not equal to what we usually mean by causality. For instance, even if  $x_1$  does not cause  $x_2$ , it may still help to predict  $x_2$ , and thus Granger-causes  $x_2$  if changes in  $x_1$  precedes that of  $x_2$  for some reason.
- A naive example is that we observe that a dragonfly flies much lower before a rain storm, due to the lower air pressure. We know that dragonflies do not cause a rain storm, but it does help to predict a rain storm, thus Granger-causes a rain storm.